# A CONTINUOUS ANALOGUE OF LATTICE PATH ENUMERATION: PART II 

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#### Abstract

Following the work of Cano and Díaz, we study continuous binomial coefficients and Catalan numbers. We explore their analytic properties, including integral identities and generalizations of discrete convolutions. We also conduct an in-depth analysis of a continuous analogue of the binomial distribution, including a stochastic representation as a Goldstein-Kac process.


## 1. Introduction

In two recent papers [1, 2], Leonardo Cano and Rafael Díaz introduced continuous analogues of the binomial coefficients and Catalan numbers. They did this by introducing a general procedure for studying continuous lattice paths, then measuring the volume of a moduli space associated to these continuous paths. By recognizing the binomial coefficients and Catalan numbers as counting certain types of lattice paths, their continuous analogues are defined as the volumes of associated moduli spaces.

In Part I of this work [10], we focused on the geometric definitions behind continuous lattice path enumeration. Our most telling result is that through a limiting procedure with Todd operators, we are able to turn results about continuous Catalan numbers into results about discrete Catalan numbers. Therefore, studying the continuous case will lead to new insight about the discrete case. In this current paper, we therefore ignore the geometric intuition underlying the continuous binomial coefficients and Catalan numbers and treat them as analytic objects of independent interest.

We already have the fundamental result:
Theorem 1. [2, Theorem 14] For $0 \leq s \leq x$, the continuous binomial coefficient $\left\{\begin{array}{l}x \\ s\end{array}\right\}$ satisfies

$$
\left\{\begin{array}{l}
x  \tag{1.1}\\
s
\end{array}\right\}=2 I_{0}(2 \sqrt{s(x-s)})+\frac{x}{\sqrt{s(x-s)}} I_{1}(2 \sqrt{s(x-s)}),
$$

where $I_{v}(z)$ denotes the modified Bessel function of the first kind.
We prove the following closed form expression for continuous Catalan numbers in Section 2.

Theorem 2. The continuous Catalan numbers defined in [2] are equal to

$$
\begin{equation*}
C(x, y)=I_{0}\left(\sqrt{x^{2}-y^{2}}\right)-\frac{x-y}{x+y} I_{2}\left(\sqrt{x^{2}-y^{2}}\right) \tag{1.2}
\end{equation*}
$$

We can regard these expressions as definitions for both objects, and indeed they lead to easy analytic continuations. The vast literature surrounding Bessel functions then means that we can prove several deep results about these two quantities, which should translate into new intuition about the discrete cases. We prove analogues of several discrete identities, such as the Chu-Vandermonde identity or Catalan convolution, and collect some integral transforms associated with both objects.

Moreover, we can naturally define, as in [2], the continuous binomial distribution $C B(x, p)$ associated to continuous binomials. It has parameters $x \geq 0$ and $0 \leq p \leq 1$, and density function

$$
f_{x, p}(s):= \begin{cases}\frac{1}{A_{x, p}}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} p^{s}(1-p)^{x-s}, & 0 \leq s \leq x  \tag{1.3}\\
0, & \text { otherwise }\end{cases}
$$

with $s \in[0, x]$. The normalization constant $A_{x, p}$ is such that

$$
\int_{0}^{x} f_{x, p}(s) d s=1
$$

and its value is given in Theorem 16.
Sections 3 and 4 lead to several convolution identities and integral transforms for the continuous binomial coefficient, along with closed form expression for the normalization constant $A_{x, p}$ and the moment generating function for $f_{x, p}(s)$. Finally, in Section 5 we are able to prove a probabilistic interpretation of the continuous binomial coefficient due to its close connection to the Goldstein-Kac telegraph process.

## 2. Continuous Catalan Numbers

2.1. Closed form. Recall that the discrete Catalan numbers $C_{n}$ count the number of lattice paths joining the points $(0,0)$ and $(n, 0)$ that stay above the $x$-axis. Therefore, the continuous Catalan numbers must satisfy similar restrictions - they correspond to continuous analogues of lattice paths that stay above the $x$ axis.

Let us first define the following polytope, which contains all possible paths in the plane made out of $n$ steps of arbitrary lengths in the East or North directions that connect the origin to the point $\left(\frac{x+y}{2}, \frac{x-y}{2}\right)$ and remain under the line $y=x$.

Definition 3. For $n \geq 1$, the convex polytope $\Lambda^{n}(x, y)$ is defined as the set of all $\left(x_{1}, \ldots, x_{n}, y_{1}, \ldots, y_{n}\right) \in \mathbb{R}^{2 n}$ that satisfy the following inequalities:

$$
\begin{aligned}
0 \leq x_{1} \leq \cdots \leq x_{n} & \leq \frac{x+y}{2} \\
0 \leq y_{1} \leq \cdots \leq y_{n} & \leq \frac{x-y}{2} \\
y_{i} & \leq x_{i} .
\end{aligned}
$$

This polytope allows to define the continuous Catalan numbers as follows.
Definition 4. [2, Defn. 23] The continuous Catalan numbers are defined by

$$
\begin{equation*}
C(x, y):=\sum_{n=0}^{\infty} \operatorname{vol}\left(\Lambda^{n}(x, y)\right), 0 \leq y \leq x \tag{2.1}
\end{equation*}
$$

where the volume is computed with respect to the Lebesgue measure.
The volume of each of these polytopes can then be explicitly computed as follows.
Lemma 5. For $n \geq 0$, the volume of $\Lambda^{n}(x, y)$ is equal to

$$
\begin{equation*}
\operatorname{vol}\left(\Lambda^{n}(x, y)\right)=\frac{(x-y)^{n}(x+y)^{n-1}(x+(2 n+1) y)}{2^{2 n} n!(n+1)!} \tag{2.2}
\end{equation*}
$$

Proof. The proof of this lemma is elementary, and follows by induction on $n$. Namely, it is easily checked that the right-hand side satisfies the integral recurrrence [2, Prop. 27])

$$
\operatorname{vol}\left(\Lambda^{n+1}(x, y)\right)=\int_{0}^{\frac{x-y}{2}} \int_{0}^{\frac{x+y}{2}-b} \operatorname{vol}\left(\Lambda^{n}(a+2 b, a)\right) d a d b
$$

together with the initial condition $\operatorname{vol}\left(\Lambda^{0}(x, y)\right)=1$.
The proof of Theorem 2 is now completed by computing the sum in (2.1) using the expression (2.2).

Moreover, by summing (2.2) over all $n$, it is shown in [2, Prop. 29] that the continuous Catalan numbers obey the recursion

$$
\begin{equation*}
C(x, y)=1+\int_{0}^{\frac{x-y}{2}} \int_{0}^{\frac{x+y}{2}-b} C(a+2 b, a) d a d b \tag{2.3}
\end{equation*}
$$

As a check, we can manually verify that the closed form (2.1) for the continuous Catalan numbers obeys the same recursion.
2.2. Parallels between the continuous and discrete case. The special case $y=0$ gives the continuous Catalan function as defined in [2]:

$$
C(2 x, 0)=\frac{I_{1}(2 x)}{x} .
$$

With $C_{n}$ denoting the usual Catalan numbers, we observe that

$$
\frac{I_{1}(2 x)}{x}=\sum_{n \geq 0} \frac{x^{2 n}}{n!(n+1)!}=\sum_{n \geq 0} \frac{x^{2 n}}{2 n!} C_{n} .
$$

Therefore, the continuous Catalan function $C(2 x, 0)$ is related to the generating function of Catalan numbers by

$$
\begin{equation*}
\int_{0}^{+\infty} C(2 \sqrt{x} u, 0) e^{-u} d u=\sum_{n} C_{n} x^{n}=\frac{2}{1+\sqrt{1-4 x}} \tag{2.4}
\end{equation*}
$$

The fact that a discrete generating function of the Catalan numbers is related to a continuous integral transform of the continuous Catalan function should not come as a surprise. However, the fact that the continuous Catalan numbers have a simple closed form expression in terms of Bessel functions lends hope to discovering closed form expressions for the continuous analogues of other objects that count lattice paths, such as the Delannoy numbers.

A further parallel between the classical and continuous cases is provided by considering the convolution identity

$$
\begin{equation*}
\sum_{k=0}^{n-1} C_{k} C_{n-k}=\Delta C_{n} \tag{2.5}
\end{equation*}
$$

with $\Delta C_{n}=C_{n+1}-C_{n}$, a consequence of the fact that the generating function

$$
c(z)=\sum_{n \geq 0} C_{n} z^{n}=\frac{2}{1+\sqrt{1-4 z}}
$$

satisfies the equation

$$
c^{2}(z)=\frac{c(z)-1}{z}
$$

For the continuous Catalan numbers, we have the following result, which is clearly a continuous analogue of the discrete identity (2.5).
Theorem 6. The continuous Catalan numbers $C(x)=C(x, 0)$ satisfy the convolution identity

$$
\begin{equation*}
\int_{0}^{z} C(x) C(z-x) d x=4 \frac{d}{d z} C(z) \tag{2.6}
\end{equation*}
$$

Proof. Since

$$
C(x)=C(x, 0)=2 \frac{I_{1}(x)}{x},
$$

the continuous equivalent of this generating function is the Laplace transform

$$
\mathcal{L}_{C}(s)=\int_{0}^{+\infty} C(x) e^{-x s} d x=\frac{2}{s+\sqrt{s^{2}-1}}
$$

Since the derivative of the continuous Catalan number

$$
C^{\prime}(x)=\frac{d}{d x} C(x)=2 \frac{I_{2}(x)}{x}
$$

has Laplace transform

$$
\mathcal{L}_{C^{\prime}}(s)=-1+2 s^{2}-2 s \sqrt{s^{2}-1}
$$

we deduce the identity

$$
\mathcal{L}_{C}^{2}(s)=4 \mathcal{L}_{C^{\prime}}(s)
$$

Taking the inverse Laplace transform of this identity gives the desired identity.
2.3. Integral representations. We calculate some useful integral representations for the continuous Catalan numbers, which enable the easy application of Laplacetransformation type proofs. These also allow us to view the continuous Catalan numbers as probability distribution functions, and recover various moment expressions for the discrete Catalan numbers.

Theorem 7. We have the integral representations

$$
C(x, y)=\frac{1}{\pi} \int_{0}^{\pi} e^{x \cos t}\left[\cos (y \sin t)-\left(\frac{x-y}{x+y}\right)^{2} \cos (y \sin t-2 t)\right] d t
$$

and

$$
C(2 x, 0)=\frac{2}{\pi} \int_{0}^{\pi} e^{2 x \cos t} \sin ^{2} t d t=\frac{I_{1}(2 x)}{x}
$$

Proof. This follows from a straightforward application of the generalized Schläfli formula [6, p. 81],

$$
\begin{aligned}
\left(\frac{a-b}{a+b}\right)^{-\frac{v}{2}} J_{v}\left(\sqrt{a^{2}-b^{2}}\right) & =\frac{1}{\pi} \int_{0}^{\pi} e^{b \cos t} \cos (a \sin t-v t) d t \\
& -\frac{\sin (\pi v)}{\pi} \int_{0}^{+\infty} e^{-a \sinh t-b \cosh t-v t} d t
\end{aligned}
$$

This is transfered to the Bessel $I$ functions using $I_{v}(z)=e^{-l v \frac{\pi}{2}} J_{v}(v z)$.
For $v=0$ we have

$$
I_{0}\left(\sqrt{a^{2}-b^{2}}\right)=\frac{1}{\pi} \int_{0}^{\pi} e^{a \cos t} \cos (b \sin t) d t
$$

for $v=1$,

$$
I_{1}\left(\sqrt{a^{2}-b^{2}}\right)=\sqrt{\frac{a-b}{a+b}} \frac{1}{\pi} \int_{0}^{\pi} e^{a \cos t} \cos (b \sin t-t) d t
$$

and for $v=2$,

$$
I_{2}\left(\sqrt{a^{2}-b^{2}}\right)=\frac{a-b}{a+b} \frac{1}{\pi} \int_{0}^{\pi} e^{a \cos t} \cos (b \sin t-2 t) d t
$$

Substituting these into the closed form expression from Theorem 2 completes the proof.

From this integral representation, we easily recover the expression (2.4)

$$
\begin{aligned}
\int_{0}^{+\infty} e^{-u} C(2 u \sqrt{x}, 0) d u & =\int_{0}^{+\infty} e^{-u} \frac{2}{\pi} \int_{0}^{\pi} e^{2 u \sqrt{x} \cos t} \sin ^{2} t d t d u \\
& =\frac{2}{\pi} \int_{0}^{\pi} \sin ^{2} t \int_{0}^{+\infty} e^{-u} e^{2 u \sqrt{x} \cos t} d u d t \\
& =\frac{2}{\pi} \int_{0}^{\pi} \frac{\sin ^{2} t}{1-2 \sqrt{x} \cos t} d t=\frac{2}{1+\sqrt{1-4 x}}
\end{aligned}
$$

The change of variable $z=\cos t$ in the second integral representation in Theorem 7 also gives

$$
C(x, 0)=\frac{2}{\pi} \int_{-1}^{+1} e^{x z} \sqrt{1-z^{2}} d z
$$

which can be expressed as

$$
C(x, 0)=\mathbb{E} e^{Z x}
$$

the moment generating function of a random variable $Z$ distributed according to the semi-circle distribution $\frac{2}{\pi} \sqrt{1-x^{2}}$. This is the continuous equivalent of the representation of Catalan numbers as the moments of the same distribution,

$$
C_{n}=\frac{2}{\pi} \int_{-1}^{1}(2 z)^{2 n} \sqrt{1-z^{2}} d z
$$

We now prove a general integral formula involving $C(x, y)$. An analogue of this formula for continuous binomial coefficients is the key element in our analysis of the continuous binomial distribution.

Theorem 8. Given any function $\Phi(y)$ supported on $[0, x]$, the integral

$$
I_{\Phi}(x)=\int_{0}^{x} C(x, y) \Phi(y) d y
$$

can be computed as

$$
I_{\Phi}(x)=\frac{1}{\pi} \int_{-1}^{+1} e^{-x u} \hat{\Phi}(u)\left(\sqrt{1+\frac{1}{u}}-1\right) \frac{d u}{\sqrt{1-u^{2}}}
$$

where

$$
\hat{\Phi}(t)=\int_{0}^{x} \Phi(y) e^{-i y t} d y
$$

is the Fourier transform of $\Phi$.
Proof. Consider a function $\Phi(y)$ with support $[0, x]$ and the integral

$$
I_{\Phi}(x)=\int_{0}^{x} C(x, y) \Phi(y) d y
$$

Then the Laplace transform $\tilde{I}_{\Phi}(p)$ of $I_{\Phi}$ can be computed as

$$
\begin{aligned}
\tilde{I}_{\Phi}(p) & =\int_{0}^{+\infty} I_{\Phi}(x) e^{-p x} d x=\int_{0}^{+\infty} \int_{0}^{x} C(x, y) \Phi(y) d y e^{-p x} d x \\
& =\int_{0}^{+\infty} \Phi(y) \int_{y}^{+\infty} e^{-p x} C(x, y) d x d y
\end{aligned}
$$

We then exploit the Laplace transforms [9, 3.15.4.2, 3.15.4.9]

$$
\begin{aligned}
\int_{y}^{+\infty} e^{-p x} I_{0}\left(\sqrt{x^{2}-y^{2}}\right) d x & =\frac{1}{\sqrt{p^{2}-1}} e^{-y \sqrt{p^{2}-1}}, \\
\int_{y}^{+\infty} e^{-p x} \frac{x-y}{x+y} I_{2}\left(\sqrt{x^{2}-y^{2}}\right) d x & =\int_{y}^{+\infty} e^{-p x} \frac{x^{2}-y^{2}}{(x+y)^{2}} I_{2}\left(\sqrt{x^{2}-y^{2}}\right) d x \\
& =\frac{1}{\left(p+\sqrt{p^{2}-1}\right)^{2}} \frac{1}{\sqrt{p^{2}-1}} e^{-y \sqrt{p^{2}-1}}
\end{aligned}
$$

to deduce

$$
\begin{aligned}
\tilde{I}_{\Phi}(p) & =\int_{0}^{+\infty} \Phi(y) \frac{1}{\sqrt{p^{2}-1}} e^{-y \sqrt{p^{2}-1}}\left[1+\frac{1}{\left(p+\sqrt{p^{2}-1}\right)^{2}}\right] d y \\
& =\frac{2 p}{p+\sqrt{p^{2}-1}} \frac{1}{\sqrt{p^{2}-1}} \tilde{\Phi}\left(\sqrt{p^{2}-1}\right),
\end{aligned}
$$

with $\tilde{\Phi}$ the Laplace transform of $\Phi$. Since

$$
\sum_{n \geq 0} \frac{C_{n}}{(4 p)^{n}}=\frac{2 p}{p+\sqrt{p^{2}-1}},
$$

we can write

$$
\tilde{I}_{\Phi}(p)=\sum_{n \geq 0} \frac{C_{n}}{2^{2 n}} \frac{1}{p^{n}} \frac{\tilde{( }\left(\sqrt{p^{2}-1}\right)}{\sqrt{p^{2}-1}}
$$

and the corresponding inverse Laplace transform

$$
I_{\Phi}(x)=\sum_{n \geq 0} \frac{C_{n}}{22^{2 n}} \mathcal{L}^{-1}\left[\frac{1}{p^{n}} \frac{\tilde{( }\left(\sqrt{p^{2}-1}\right)}{\sqrt{p^{2}-1}}\right] .
$$

Now we can use the results

$$
\mathcal{L}^{-1}\left[\frac{\tilde{\Phi}\left(\sqrt{p^{2}-1}\right)}{\sqrt{p^{2}-1}}\right]=\int_{0}^{x} I_{0}\left(\sqrt{x^{2}-y^{2}}\right) \Phi(y) d y
$$

and

$$
\mathcal{L}^{-1}\left[\frac{F(p)}{p^{n}}\right]=f^{(-n)}(t),
$$

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where $f^{(-n)}$ is the $n-$ th antiderivative of $f$. Starting from the integral representation

$$
I_{0}\left(\sqrt{x^{2}-y^{2}}\right)=\frac{1}{2 \pi} \int_{0}^{2 \pi} e^{-(x \cos \theta+x y \cos \theta)} d \theta
$$

we have

$$
\begin{aligned}
\int_{0}^{x} I_{0}\left(\sqrt{x^{2}-y^{2}}\right) \Phi(y) d y & =\int_{0}^{x} \frac{1}{2 \pi} \int_{0}^{2 \pi} e^{-(x \cos \theta+l y \cos \theta)} d \theta \Phi(y) d y \\
& =\frac{1}{2 \pi} \int_{0}^{2 \pi} e^{-x \cos \theta} d \theta \int_{0}^{x} \Phi(y) e^{-l y \cos \theta} d y
\end{aligned}
$$

Since the function $\Phi(y)$ has bounded support on $[0, x]$, the inner integral is recognized as its Fourier transform $\hat{\Phi}$ computed at $\cos \theta$, and

$$
\int_{0}^{x} I_{0}\left(\sqrt{x^{2}-y^{2}}\right) \Phi(y) d y=\frac{1}{2 \pi} \int_{0}^{2 \pi} e^{-x \cos \theta} \hat{\Phi}(\cos \theta) d \theta
$$

The antiderivative of order $n$ in $x$ is easily computed as

$$
\frac{(-1)^{n}}{2 \pi} \int_{0}^{2 \pi} \frac{e^{-x \cos \theta}}{(\cos \theta)^{n}} \hat{\Phi}(\cos \theta) d \theta
$$

Consequently,

$$
\begin{aligned}
I_{\Phi}(x) & =\sum_{n \geq 0} \frac{C_{n}}{2^{2 n}} \frac{(-1)^{n}}{2 \pi} \int_{0}^{2 \pi} \frac{e^{-x \cos \theta}}{(\cos \theta)^{n}} \hat{\Phi}(\cos \theta) d \theta \\
& =\frac{1}{2 \pi} \int_{0}^{2 \pi} e^{-x \cos \theta} \hat{\Phi}(\cos \theta) 2 \cos \theta\left(-1+\sqrt{1+\frac{1}{\cos \theta}}\right) d \theta \\
& =\frac{1}{\pi} \int_{-1}^{+1} e^{-x u} \hat{\Phi}(u)\left(\sqrt{1+\frac{1}{u}}-1\right) \frac{d u}{\sqrt{1-u^{2}}}
\end{aligned}
$$

which is the desired result.

## 3. Continuous Binomial Coefficients

3.1. Integrals. Here we examine some of the properties of the continuous binomials $\left\{\begin{array}{l}x \\ s\end{array}\right\}$, including some integral transforms that will allow us to prove several more general theorems in Sections 4 and 5. We start with a general integral transform that will later appear in the analysis of the continuous binomial distribution.

Theorem 9. The function

$$
J_{\Phi}(x)=\int_{0}^{x}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} \Phi(s) d s
$$

has Laplace transform

$$
\int_{0}^{+\infty} J_{\Phi}(x) e^{-p x} d x=\left(\frac{1+p}{p}\right)^{2} \tilde{\Phi}\left(p-\frac{1}{p}\right)-\tilde{\Phi}(p)
$$

where $\tilde{\Phi}(p)$ is the Laplace transform of $\Phi(s)$.
As a consequence,

$$
\begin{equation*}
J_{\Phi}(x)=\mathcal{L}^{-1}\left[\left(\frac{1+p}{p}\right)^{2} \tilde{\Phi}\left(p-\frac{1}{p}\right)\right]-\Phi(x), \tag{3.1}
\end{equation*}
$$

where $\mathcal{L}^{-1}$ denotes the inverse Laplace transform.
Proof. We apply Fubini's theorem to transform the double integral

$$
\begin{aligned}
\int_{0}^{+\infty} J_{\Phi}(x) e^{-p x} d x & =\int_{0}^{+\infty} \int_{0}^{x}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} \Phi(s) d s e^{-p x} d x \\
& =\int_{0}^{+\infty} \Phi(s) \int_{s}^{+\infty} e^{-p x}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} d x d s .
\end{aligned}
$$

The inner integral is now evaluated using the change of variable $x=s+w$ as

$$
\int_{s}^{+\infty} e^{-p x}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} d x=e^{-s p} \int_{0}^{+\infty} e^{-p w}\left\{\begin{array}{c}
s+w \\
s
\end{array}\right\} d w .
$$

Using the closed form (1) for the continuous binomial coefficient, we deduce
$\int_{0}^{+\infty} e^{-p w}\left\{\begin{array}{c}s+w \\ s\end{array}\right\} d w=2 \int_{0}^{+\infty} e^{-p w} I_{0}(2 \sqrt{s w}) d w+\int_{0}^{+\infty} e^{-p w} \frac{w+s}{\sqrt{s w}} I_{1}(2 \sqrt{s w}) d w$.
These Laplace transforms can be found in [3, 6.614.3 and 6.643.2] and evaluate to

$$
\begin{gathered}
\int_{0}^{+\infty} e^{-p w} I_{0}(2 \sqrt{s w}) d w=\frac{1}{p} e^{\frac{s}{p}}, \\
\int_{0}^{+\infty} e^{-p w} \frac{w}{\sqrt{s w}} I_{1}(2 \sqrt{s w}) d w=\frac{1}{p^{2}} e^{\frac{s}{p}},
\end{gathered}
$$

and

$$
\int_{0}^{+\infty} e^{-p w} \frac{1}{\sqrt{s w}} I_{1}(2 \sqrt{s w}) d w=-1+e^{\frac{s}{p}} .
$$

We deduce

$$
\int_{0}^{+\infty} e^{-p w}\left\{\begin{array}{c}
s+w  \tag{3.2}\\
s
\end{array}\right\} d w=e^{\frac{s}{p}} \frac{p^{2}+2 p+1}{p^{2}}-1=e^{\frac{s}{p}}\left(\frac{p+1}{p}\right)^{2}-1 .
$$

This is now substituted in the outer integral to obtain

$$
\begin{gathered}
\int_{0}^{+\infty} \Phi(s) e^{-s p}\left\{e^{\frac{s}{p}}\left(\frac{p+1}{p}\right)^{2}-1\right\} d s \\
=\left(\frac{p+1}{p}\right)^{2} \int_{0}^{+\infty} \Phi(s) e^{-s\left(p-\frac{1}{p}\right)} d s-\int_{0}^{+\infty} \Phi(s) e^{-s p} d s .
\end{gathered}
$$

These two integral are recognized as the Laplace transforms of $\Phi(s)$ computed respectively at $p-\frac{1}{p}$ and $p$, and the result follows.

There are several important special cases.
Corollary 10. Choosing $\Phi(s)=\alpha^{s} e^{u s}$, we deduce the value of the integral

$$
\begin{align*}
\int_{0}^{x}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} \alpha^{s} e^{u s} d s & =2 \alpha^{\frac{x}{2}} e^{\frac{u x}{2}}\left\{\cosh \left(\frac{x}{2} \sqrt{4+(u+\log \alpha)^{2}}\right)-\cosh \left(\frac{x}{2}(u+\log \alpha)\right)\right.  \tag{3.3}\\
& \left.+\frac{2}{\sqrt{4+(u+\log \alpha)^{2}}} \sinh \left(\frac{x}{2} \sqrt{4+(u+\log \alpha)^{2}}\right)\right\} .
\end{align*}
$$

Proof. Choose $\Phi(s)=\alpha^{s} e^{u s}$ so that $\tilde{\Phi}(p)=\frac{1}{p-u-\log \alpha}$ and use formula (3.1) to obtain the result.

Another consequence is as follows.
Corollary 11. The integral

$$
\int_{0}^{x}\left\{\begin{array}{l}
x  \tag{3.4}\\
s
\end{array}\right\} d s=2\left(e^{x}-1\right)
$$

holds, as computed in [2].
3.2. Continuous Chu-Vandermonde formula. Now that we have obtained continuous binomial coefficients with nice reductions to the discrete case, we can try to find continuous generalizations of discrete identities. We first consider an averaged case of the Chu-Vandermonde identity,

$$
\sum_{k}\binom{k+s_{1}}{s_{1}}\binom{n-k+s_{2}}{s_{2}}=\binom{n+s_{1}+s_{2}+1}{n}
$$

which can be regarded as the prototypical binomial convolution. We denote the Dirac delta distribution by $\delta(x)$ and by $*$ the integral convolution

$$
f * g(x)=\int f(u) g(x-u) d u
$$

and notice that $f * \delta=f$.
Theorem 12. The continuous binomial coefficient satisfies the identity

$$
\begin{gathered}
\left(\delta(x)+\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\}\right) *\left(\delta(x)+\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\}\right) \\
=\left(\delta(x)+\left\{\begin{array}{c}
x+\frac{s_{1}+s_{2}}{2} \\
\frac{s_{1}+s_{2}}{2}
\end{array}\right\}\right) *\left(\delta(x)+\left\{\begin{array}{c}
x+\frac{s_{1}+s_{2}}{2} \\
\frac{s_{1}+s_{2}}{2}
\end{array}\right\}\right)
\end{gathered}
$$

to be compared to the discrete version

$$
\sum_{k}\binom{k+s_{1}}{s_{1}}\binom{n-k+s_{2}}{s_{2}}=\sum_{k}\binom{k+\frac{s_{1}+s_{2}}{2}}{\frac{s_{1}+s_{2}}{2}}\binom{n-k+\frac{s_{1}+s_{2}}{2}}{\frac{s_{1}+s_{2}}{2}}
$$

Proof. From (3.2), the Laplace transform (in the variable $x$ ) of the continuous binomial coefficient

$$
f_{s}(x)=\left\{\begin{array}{c}
x+s \\
s
\end{array}\right\}
$$

is

$$
F_{s}(p)=\int_{0}^{+\infty} e^{-x p}\left\{\begin{array}{c}
x+s \\
s
\end{array}\right\} d x=e^{\frac{s}{p}}\left(\frac{p+1}{p}\right)^{2}-1
$$

We deduce

$$
\begin{aligned}
F_{s_{1}}(p) F_{s_{2}}(p) & =\left(e^{\frac{s_{1}}{p}}\left(\frac{p+1}{p}\right)^{2}-1\right)\left(e^{\frac{s_{2}}{p}}\left(\frac{p+1}{p}\right)^{2}-1\right) \\
& =e^{\frac{s_{1}+s_{2}}{p}}\left(\frac{p+1}{p}\right)^{4}-1-\left(e^{\frac{s_{1}}{p}}\left(\frac{p+1}{p}\right)^{2}-1\right)-\left(e^{\frac{s_{2}}{p}}\left(\frac{p+1}{p}\right)^{2}-1\right) .
\end{aligned}
$$

The decomposition

$$
e^{\frac{s_{1}+s_{2}}{p}}\left(\frac{p+1}{p}\right)^{4}-1=\left(e^{\frac{s_{1}+s_{2}}{2 p}}\left(\frac{p+1}{p}\right)^{2}-1\right)\left(\left(e^{\frac{s_{1}+s_{2}}{2 p}}\left(\frac{p+1}{p}\right)^{2}-1\right)+2\right)
$$

gives

$$
F_{s_{1}}(p) F_{s_{2}}(p)=F_{\frac{s_{1}+s_{2}}{2}}(p)\left(F_{\frac{s_{1}+s_{2}}{2}}(p)+2\right)-F_{s_{1}}(p)-F_{s_{2}}(p)
$$

the inverse Laplace transform of which is

$$
\left(f_{s_{1}} * f_{s_{2}}\right)(x)=\left(f_{\frac{s_{1}+s_{2}}{2}} * f_{\frac{s_{1}+s_{2}}{2}}\right)(x)+2 f_{\frac{s_{1}+s_{2}}{2}}(x)-f_{s_{1}}(x)-f_{s_{2}}(x)
$$

Equivalently,

$$
\begin{aligned}
& \left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\} *\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\}+\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\}+\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\} \\
& =\left\{\begin{array}{c}
x+\frac{s_{1}+s_{2}}{2} \\
\frac{s_{1}+s_{2}}{2}
\end{array}\right\} *\left\{\begin{array}{c}
x+\frac{s_{1}+s_{2}}{2} \\
\frac{s_{1}+s_{2}}{2}
\end{array}\right\}+2\left\{\begin{array}{c}
x+\frac{s_{1}+s_{2}}{2} \\
\frac{s_{1}+s_{2}}{2}
\end{array}\right\} .
\end{aligned}
$$

The theorem follows after rewriting this identity in terms of Dirac delta functions.

We can then give an analogue of the Chu-Vandermonde identity, based on discrete difference and differential operators. We begin with the discrete case: define the $\star$ operator as

$$
\begin{equation*}
\binom{n+k_{1}}{k_{1}} \star\binom{n+k_{2}}{k_{2}}=\sum_{m=1}^{n-1}\binom{m+k_{1}}{k_{1}}\binom{n-m+k_{2}}{k_{2}}, \tag{3.5}
\end{equation*}
$$

so that the Chu-Vandermonde identity reads:

$$
\begin{aligned}
& \binom{n+k_{1}}{k_{1}} \star\binom{n+k_{2}}{k_{2}}+\binom{n+k_{1}}{k_{1}}+\binom{n+k_{2}}{k_{2}} \\
= & \binom{n+k_{1}+k_{2}+1}{k_{1}+k_{2}+1}=\left(1+\Delta_{k_{1}+k_{2}}\right)\binom{n+k_{1}+k_{2}}{k_{1}+k_{2}},
\end{aligned}
$$

where $\Delta_{k}$ is the forward discrete difference operator in the variable $k$,

$$
\Delta_{k} f(n+k)=f(n+k+1)-f(n+k)
$$

Its continuous analogue is as follows.
Theorem 13. With $\bar{s}=s_{1}+s_{2}$, the continuous binomial coefficient satisfies the differential equation

$$
\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\} *\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\}+\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\}+\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\}=\left(1+\frac{\partial}{\partial \bar{s}}\right)^{2}\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\} .
$$

Proof. Applying Theorem 12, we have

$$
\begin{aligned}
& \left(\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\}+\delta(x)\right) *\left(\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\}+\delta(x)\right) \\
& =\left(\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x)\right) *\left(\left\{\begin{array}{l}
x \\
0
\end{array}\right\}+\delta(x)\right)
\end{aligned}
$$

where

$$
\left\{\begin{array}{l}
x \\
0
\end{array}\right\}=x+2
$$

is deduced from the Laplace transform

$$
\mathcal{L}\left(\left\{\begin{array}{l}
x \\
0
\end{array}\right\}+\delta(x)\right)=\left(1+\frac{1}{p}\right)^{2}=\mathcal{L}(2+x+\delta(x)) .
$$

We thus need to compute the convolution

$$
\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\} *(2+x)=\int_{0}^{x}(2+x-u)\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\} d u .
$$

First, using the differential equation

$$
\frac{\partial}{\partial u} \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\}=\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\},
$$

we deduce

$$
\begin{aligned}
\int_{0}^{x}\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\} d u & =\int_{0}^{x} \frac{\partial}{\partial u} \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\} d u=\frac{\partial}{\partial \bar{s}} \int_{0}^{x} \frac{\partial}{\partial u}\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\} d u \\
& =\frac{\partial}{\partial \bar{s}}\left(\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}-\left\{\begin{array}{c}
\bar{s} \\
\bar{s}
\end{array}\right\}\right)=\frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}-1
\end{aligned}
$$

This argument also shows that an antiderivative of $\left\{\begin{array}{c}u+\bar{s} \\ u\end{array}\right\}$ is $\frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}x+\bar{s} \\ x\end{array}\right\}$.

Next, integrating by parts gives

$$
\begin{aligned}
\int_{0}^{x} u\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\} d u & =\left[u \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}\right]_{0}^{x}-\int_{0}^{x} \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\} d u \\
& =x \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}-\frac{\partial^{2}}{\partial \bar{s}^{2}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\} .
\end{aligned}
$$

We deduce

$$
\begin{gathered}
\int_{0}^{x}(2+x-u)\left\{\begin{array}{c}
u+\bar{s} \\
u
\end{array}\right\} d u \\
=(2+x)\left(\frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}-1\right)-x \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}+\frac{\partial^{2}}{\partial \bar{s}^{2}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\} \\
=-(x+2)+2 \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}+\frac{\partial^{2}}{\partial \bar{s}^{2}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\} .
\end{gathered}
$$

Finally, we deduce the convolution

$$
\begin{aligned}
& \left(\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\}+\delta(x)\right) *\left(\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\}+\delta(x)\right) \\
& =\left(\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x)\right) *\left(\left\{\begin{array}{l}
x \\
0
\end{array}\right\}+\delta(x)\right) \\
& =\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\} *\left\{\begin{array}{l}
x \\
0
\end{array}\right\}+\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\left\{\begin{array}{c}
x \\
0
\end{array}\right\}+\delta(x) \\
& =-(x+2)+2 \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}+\frac{\partial^{2}}{\partial \bar{s}^{2}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}+\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\left\{\begin{array}{l}
x \\
0
\end{array}\right\}+\delta(x) \\
& =2 \frac{\partial}{\partial \bar{s}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}+\frac{\partial^{2}}{\partial \bar{s}^{2}}\left\{\begin{array}{c}
x+\bar{s} \\
x
\end{array}\right\}+\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x) \\
& =\left(1+\frac{\partial}{\partial \bar{s}}\right)^{2}\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x) .
\end{aligned}
$$

A more direct proof involves converting every term into the Laplace transform domain: start with

$$
\begin{aligned}
\mathcal{L}\left(\left(\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\}+\delta(x)\right) *\right. & \left.\left(\left\{\begin{array}{c}
x+s_{2} \\
s_{2}
\end{array}\right\}+\delta(x)\right)\right)=e^{\frac{s_{1}+s_{2}}{p}}\left(1+\frac{1}{p}\right)^{4} \\
& =e^{\frac{\bar{s}}{p}}\left(1+\frac{1}{p}\right)^{4}
\end{aligned}
$$

and

$$
\mathcal{L}\left(\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x)\right)=e^{\frac{\bar{s}}{p}}\left(1+\frac{1}{p}\right)^{2} .
$$

Since

$$
\frac{\partial}{\partial \bar{s}} e^{\frac{\overline{5}}{p}}=\frac{1}{p} e^{\frac{\overline{5}}{p}},
$$

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it follows that

$$
\left(1+\frac{\partial}{\partial \bar{s}}\right)^{2}\left(\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x)\right)=\left(1+\frac{\partial}{\partial \bar{s}}\right)^{2}\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x)
$$

has Laplace transform

$$
e^{\frac{\bar{s}}{p}}\left(1+\frac{1}{p}\right)^{2}\left(1+\frac{1}{p}\right)^{2}=e^{\frac{\bar{s}}{p}}\left(1+\frac{1}{p}\right)^{4}
$$

This second proof also allows us to state the following generalization
Corollary 14. With $\bar{s}=\sum_{i=1}^{p} s_{i}$, we have

$$
\left(\left\{\begin{array}{c}
x+s_{1} \\
s_{1}
\end{array}\right\}+\delta(x)\right) * \cdots *\left(\left\{\begin{array}{c}
x+s_{p} \\
s_{p}
\end{array}\right\}+\delta(x)\right)=\left(1+\frac{\partial}{\partial \bar{s}}\right)^{2 p}\left\{\begin{array}{c}
x+\bar{s} \\
\bar{s}
\end{array}\right\}+\delta(x) .
$$

3.3. Central binomial coefficients. The central binomial coefficients $\left\{\begin{array}{c}2 s \\ s\end{array}\right\}$ have explicit expression

$$
\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\}=2 I_{0}(2 s)+2 I_{1}(2 s)=\left(2+\frac{d}{d s}\right) I_{0}(2 s)
$$

and Laplace transform

$$
\mathcal{L}\left(\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\}+\delta(s)\right)=\sqrt{\frac{p+2}{p-2}} .
$$

The parallel with the usual central binomial coefficients already appears in the asymptotic behavior: as it is well known, for large $n$,

$$
\frac{1}{2^{2 n}}\binom{2 n}{n} \sim \frac{1}{\sqrt{\pi n}}
$$

whereas elementary asymptotic behavior results on Bessel I functions give, for large values of $s$,

$$
\frac{1}{2} e^{-4 s}\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\} \sim \frac{1}{\sqrt{\pi s}}
$$

The next theorem gives the continuous analogue of the convolution identity for central binomial coefficients

$$
\sum_{k=0}^{n}\binom{2 k}{k}\binom{2 n-2 k}{n-k}=4^{n}
$$

that can be deduced from the Taylor series

$$
\sum_{n \geq 0}\binom{2 n}{n} z^{n}=\frac{1}{\sqrt{1-4 z}}
$$

To make this analogue clearer, let us first rewrite this identity in terms of the $\star$ operator (3.5) as

$$
\binom{2 n}{n} \star\binom{2 n}{n}=4^{n}-2\binom{2 n}{n} .
$$

Theorem 15. The convolution of continuous central binomial coefficients is given by

$$
\left(\left\{\begin{array}{c}
2 s  \tag{3.6}\\
s
\end{array}\right\}+\delta(s)\right) *\left(\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\}+\delta(s)\right)=4 e^{2 s}+\delta(s),
$$

or equivalently by

$$
\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\} *\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\}=4 e^{2 s}-2\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\}
$$

This can be generalized to any $2 n$-tuple convolution as

$$
\left(\left\{\begin{array}{c}
2 s \\
s
\end{array}\right\}+\delta(s)\right)^{* 2 n}=4 e^{2 s} L_{n-1}^{(1)}(-4 s)+\delta(s)
$$

where $L_{n}^{(k)}(x)$ is the associated Laguerre polynomial with Rodrigues formula

$$
L_{n}^{(k)}(x)=\frac{e^{x} x^{-k}}{n!} \frac{d^{n}}{d x^{n}}\left(e^{-x} x^{n+k}\right)
$$

Proof. Expanding

$$
\left(\sqrt{\frac{p+2}{p-2}}\right)^{2 n}=\left(1+\frac{4}{p-2}\right)^{n}=\sum_{m=0}^{n}\binom{n}{m}\left(\frac{4}{p-2}\right)^{m}
$$

produces the inverse Laplace transform

$$
\begin{aligned}
\sum_{m=1}^{n}\binom{n}{m} \frac{2^{2 m-3} e^{2 s} s^{m-1}}{3(m-1)!}+\delta(s) & =\delta(s)+e^{2 s} \sum_{m=0}^{n-1}\binom{n}{m+1} 4.2^{2 m} \frac{s^{p}}{p!} \\
& =\delta(s)+4 e^{2 s} L_{n-1}^{(1)}(-4 s)
\end{aligned}
$$

## 4. The continuous binomial distribution

Following Cano and Díaz [2], the continuous binomial coefficients allow us to define a continuous version of the discrete binomial distribution through the probability density function

$$
f_{x, p}(s):= \begin{cases}\frac{1}{A_{x, p}}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} p^{s}(1-p)^{x-s}, & 0 \leq s \leq x  \tag{4.1}\\
0, & \text { otherwise }\end{cases}
$$

where $0 \leq p \leq 1$ and the normalization constant $A_{x, p}$ is such that

$$
\int_{0}^{x} f_{x, p}(s) d s=1
$$

Notice that the centered version of this distribution, namely the distribution of the shifted random variable

$$
Y=X-\frac{x}{2}
$$

where $X$ is distributed as in (4.1), is studied in [2]. Its density is

$$
f_{x, p}(s):= \begin{cases}\frac{1}{A_{x, p}}\left\{\begin{array}{c}
x \\
\frac{x}{2}+s
\end{array}\right\} p^{s+\frac{x}{2}}(1-p)^{\frac{x}{2}-s}, & -\frac{x}{2} \leq s \leq \frac{x}{2}  \tag{4.2}\\
0, & \text { otherwise }\end{cases}
$$

The normalization constant $A_{x, p}$ of this density is not evaluated in [2]; we give its value as follows.

Theorem 16. The normalization constant of the continuous binomial distribution is equal to

$$
\begin{aligned}
A_{x, p} & =2[p(1-p)]^{\frac{x}{2}}\left\{\cosh \left(\frac{x}{2} \sqrt{4+\log ^{2} \frac{p}{1-p}}\right)-\cosh \left(\frac{x}{2} \log \frac{p}{1-p}\right)\right. \\
& \left.+\frac{2}{\sqrt{4+\log ^{2} \frac{p}{1-p}}} \sinh \left(\frac{x}{2} \sqrt{4+\log ^{2} \frac{p}{1-p}}\right)\right\} .
\end{aligned}
$$

Proof. Since

$$
A_{x, p}=\int_{0}^{x}\left\{\begin{array}{l}
x \\
s
\end{array}\right\} p^{s}(1-p)^{x-s} d s=(1-p)^{x} \int_{0}^{x}\left\{\begin{array}{l}
x \\
s
\end{array}\right\}\left(\frac{p}{1-p}\right)^{s} d s
$$

using (3.3) with $\alpha=\frac{p}{1-p}$ and $u=0$ yields the result.
Moreover, the moment generating function of the continuous binomial distribution (4.1) can be computed explicitly as follows.

Theorem 17. The moment generating function of the continuous binomial distribution (4.1) is

$$
\begin{equation*}
\mathbb{E} e^{u X}=e^{\frac{u x}{2}} \frac{\varphi_{x, p}(u)}{\varphi_{x, p}(0)} \tag{4.3}
\end{equation*}
$$

with

$$
\begin{align*}
\varphi_{x, p}(u) & =\cosh \left(\frac{x}{2} \sqrt{4+\left(u+\log \frac{p}{1-p}\right)^{2}}\right)-\cosh \left(\frac{x}{2} \log \frac{p}{1-p}\right)  \tag{4.4}\\
& +\frac{2}{\sqrt{4+\left(u+\log \frac{p}{1-p}\right)^{2}}} \sinh \left(\frac{x}{2} \sqrt{4+\left(u+\log \frac{p}{1-p}\right)^{2}}\right) .
\end{align*}
$$

Proof. By definition,

$$
\mathbb{E} e^{u X}=\int_{0}^{x} e^{u s} f_{x, p}(s) d s=\frac{(1-p)^{x}}{A_{x, p}} \int_{0}^{x} e^{u s}\left\{\begin{array}{l}
x  \tag{4.5}\\
s
\end{array}\right\}\left(\frac{p}{1-p}\right)^{s} d s
$$

Applying (3.3) yields

$$
\begin{gathered}
\mathbb{E} e^{u X}=\frac{(1-p)^{x}}{A_{x, p}} 2\left(\frac{p}{1-p}\right)^{\frac{x}{2}} e^{\frac{u x}{2}}\left\{\cosh \left(\frac{x}{2} \sqrt{4+\left(u+\log \frac{p}{1-p}\right)^{2}}\right)\right. \\
\left.-\cosh \left(\frac{x}{2} \log \frac{p}{1-p}\right)+\frac{2}{\sqrt{4+\left(u+\log \frac{p}{1-p}\right)^{2}}} \sinh \left(\frac{x}{2} \sqrt{4+\left(u+\log \frac{p}{1-p}\right)^{2}}\right)\right\} \\
=\frac{2[p(1-p)]^{\frac{x}{2}}}{A_{x, p}} e^{\frac{u x}{2}} \varphi_{x, p}(u)
\end{gathered}
$$

with $\varphi_{x, p}(u)$ defined by (4.4). Taking $u=0$ yields $1=\frac{2[p(1-p)]^{\frac{x}{2}}}{A_{x, p}} \varphi_{x, p}(0)$, from which (4.3) follows.

We remark that $\varphi_{x, p}(z)=\varphi_{x, \frac{1}{2}}\left(z+\log \frac{p}{1-p}\right)$.
In the symmetric case $p=\frac{1}{2}$, the moments can be explicitly computed, following the approach used by S.M. Iacus and N. Yoshida in the case of the telegraph process [4].

Theorem 18. The moments of a random variable $X$ distributed according to the symmetric discrete binomial distribution (4.2) density with $p=\frac{1}{2}$ are

$$
\mathbb{E} X^{k}= \begin{cases}\frac{1}{\left(e^{x}-1\right)}\left[\left(\frac{x}{2}\right)^{\frac{k+1}{2}} \Gamma\left(\frac{k+1}{2}\right)\left(I_{\frac{k+1}{2}}(x)+I_{\frac{k-1}{2}}(x)\right)-\left(\frac{x}{2}\right)^{k}\right], & \text { keven } \\ 0, & k \text { odd } .\end{cases}
$$

Proof. The density in the case $p=\frac{1}{2}$ is

$$
f_{x}(s)=\frac{1}{2\left(e^{x}-1\right)}\left\{\begin{array}{c}
x \\
\frac{x}{2}+s
\end{array}\right\}
$$

so that

$$
\begin{aligned}
\mathbb{E} X^{k} & =\int_{-\frac{x}{2}}^{+\frac{x}{2}} s^{k} f_{x}(s) d s=\frac{1}{2\left(e^{x}-1\right)} \int_{-\frac{x}{2}}^{+\frac{x}{2}} s^{k}\left[2 I_{0}\left(2 \sqrt{\frac{x^{2}}{4}-s^{2}}\right)\right] d s \\
& +\frac{1}{2\left(e^{x}-1\right)} \int_{-\frac{x}{2}}^{+\frac{x}{2}} s^{k}\left[\frac{x}{\sqrt{\frac{x^{2}}{4}-s^{2}}} I_{1}\left(2 \sqrt{\frac{x^{2}}{4}-s^{2}}\right)\right] d s .
\end{aligned}
$$

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Figure 4.1. The continuous binomial distribution for $p=\frac{1}{2}, \frac{1}{3}$, and $\frac{1}{4}$, right to left.

The first integral

$$
\int_{-\frac{x}{2}}^{+\frac{x}{2}} s^{k}\left[2 I_{0}\left(2 \sqrt{\frac{x^{2}}{4}-s^{2}}\right)\right] d s=\left(1+(-1)^{k}\right)\left(\frac{x}{2}\right)^{\frac{k+1}{2}} \Gamma\left(\frac{k+1}{2}\right) I_{\frac{k+1}{2}}(x)
$$

is computed in [4] by considering the Taylor expansion of the the Bessel function, but it can also be deduced from Entry 2.15.2.6 in [8] after the change of variable $z=\sqrt{\frac{x^{2}}{4}}-s^{2}$. The second integral

$$
\begin{gathered}
x \int_{-\frac{x}{2}}^{+\frac{x}{2}} s^{k}\left[\frac{1}{\sqrt{\frac{x^{2}}{4}-s^{2}}} I_{1}\left(2 \sqrt{\frac{x^{2}}{4}-s^{2}}\right)\right] d s \\
=\left(1+(-1)^{k}\right)\left\{\left(\frac{x}{2}\right)^{\frac{k+1}{2}} \Gamma\left(\frac{k+1}{2}\right) I_{\frac{k-1}{2}}(x)-\left(\frac{x}{2}\right)^{k}\right\},
\end{gathered}
$$

is also computed in [4] using the same technique, and does not seem to appear in the usual tables of integrals. Using these two integrals yields the result.

The continuous binomial distribution is illustrated for $x=10$ and for the 3 values $p=\frac{1}{2}$ (symmetric curve), $p=\frac{1}{3}$ and $p=\frac{1}{4}$.

## 5. A Stochastic Representation

The continuous binomial coefficient can be related to a stochastic process, the Goldstein-Kac telegraph process. This was studied by E. Orsingher in [7] and a complete introduction to this process is given in [5]. The Goldstein-Kac process describes successive changes of a binary state, the number of these changes following a Poisson distribution: this implies that the successive times spent in each statein our case, the lengths traveled in each successive direction-are independently


Figure 5.1. A trajectory of a telegraph process
and uniformly distributed. This corresponds to the least informative (maximum entropy) among all bounded support distributions.

Consider a Poisson process $n(t)$ with parameter $\lambda>0$, and a particle that travels on the real axis, starting from 0 with an initial velocity equal to $+c$ or $-c$ each with probability $1 / 2$. The velocity of the particle is supposed to be

$$
v(t)= \pm c(-1)^{n(t)}
$$

so that the particle changes instantaneously the sign of its constant velocity $c$ at each Poisson event. One trajectory of the velocity in the case $\lambda=1.3$ and $t \in[0,15]$ is given below.

The location $X(t)$ of the particle at time $t$, given by

$$
\begin{equation*}
X(t)=\int_{0}^{t} v(\tau) d \tau= \pm c \int_{0}^{t}(-1)^{n(\tau)} d \tau \tag{5.1}
\end{equation*}
$$

defines the Goldstein-Kac process. The probability function of the location of the particle at time $t$ has two parts:

- The discrete part is

$$
\operatorname{Pr}\{X(t)=c t \mid n(t)=0\}=\operatorname{Pr}\{X(t)=-c t \mid n(t)=0\}=\frac{1}{2} e^{-\lambda t}
$$

which is the conditional probability that the particle has reached position $\pm c t$ at time $t$ without any Poisson event happening since it started at time 0

- Conditionally to the event $n(t)>0$, the probability function

$$
p(s, t) d s=\operatorname{Pr}\{s \leq X(t)<s+d s\}
$$



Figure 5.2. Two trajectories induced by the integrated telegraph process
is continuous and its density is given by

$$
\begin{equation*}
p(s, t)=\frac{e^{-\lambda t}}{2 c}\left[\lambda I_{0}\left(\frac{\lambda}{c} \sqrt{c^{2} t^{2}-s^{2}}\right)+\frac{\lambda}{c} \frac{t c^{2}}{\sqrt{c^{2} t^{2}-s^{2}}} I_{1}\left(\frac{\lambda}{c} \sqrt{c^{2} t^{2}-s^{2}}\right)\right] . \tag{5.2}
\end{equation*}
$$

Now take

$$
c t=\frac{x}{2} \text { and } \lambda=2 c=\frac{x}{t}
$$

so that

$$
\begin{aligned}
p(s, t) & =\frac{e^{-x}}{2}\left[2 I_{0}\left(2 \sqrt{\frac{x^{2}}{4}-s^{2}}\right)+\frac{x}{\sqrt{\frac{x^{2}}{4}-s^{2}}} I_{1}\left(2 \sqrt{\frac{x^{2}}{4}-s^{2}}\right)\right] \\
& =\frac{e^{-x}}{2}\left\{\begin{array}{c}
x \\
\frac{x}{2}+s
\end{array}\right\} .
\end{aligned}
$$

In the discrete setup of a centered binomial distribution with $p=1-p=\frac{1}{2}$, the usual binomial coefficient $\binom{n}{k}$ is proportional to the numbers of ways that the particle, starting from 0 , can reach the site $k$ after $n$ independent equiprobable jumps to the left or to the right. Assuming $n$ even, we have $-\frac{n}{2} \leq k \leq \frac{n}{2}$.

Similarly, the continuous binomial coefficient measures the "number" of continuous paths of an integrated telegraphic random process that, starting from $(0,0)$, reach the point $\left(\frac{x}{2}+s, \frac{x}{2}-s\right)$, traveling horizontally during a total time $\frac{1}{c}\left(\frac{x}{2}+s\right)$ and vertically during a remaining total time $\frac{1}{c}\left(\frac{x}{2}-s\right)$, and switching between East and North directions each time a Poisson event happens. The attached figure shows two trajectories of such a process.

Note that the density (5.2) satisfies the differential equation

$$
\begin{equation*}
c^{2} \frac{\partial^{2} p}{\partial x^{2}}=\frac{\partial^{2} p}{\partial t^{2}}+2 \lambda \frac{\partial p}{\partial t} \tag{5.3}
\end{equation*}
$$

which can be transformed into

$$
c^{2} \frac{\partial^{2} v}{\partial x^{2}}+\lambda v^{2}=\frac{\partial^{2} v}{\partial t^{2}}
$$

with $v(x, t) e^{-\lambda t}=p(x, t)$.

## 6. Next Steps

In this work, we studied coutinuous analogs of the binomial coefficient and Catalan numbers, and showed that they possess several properties of independent interest. Compact expressions for both in terms of Bessel I functions should allow us to prove several straightforward results about them in the future. Because of a reduction procedure to the discrete case, described in [10], this can potentially inform research about the discrete case.

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## References

[1] L. Cano and R. Díaz, Indirect influences on directed manifolds, Advanced Studies in Comterporary Mathematics, 28, 93-114, 2018.
[2] L. Cano and R. Díaz, Continuous analogues for the binomial coefficients and the Catalan numbers. ArXiV:1602.09132v4[math.CO].
[3] I. S. Gradshteyn and I. M. Ryzhik, eds., Table of integrals, series, and products. 7th ed., Academic Press, San Diego, 2007.
[4] S.M. Iacus and N. Yoshida, Estimation for the discretely observed telegraph process, Theor. Probability and Math. Statist. 78, 37-47, 2009.
[5] A. D. Kolesnik, Moment analysis of the telegraph random process, Buletinul Academiei De Stiințe a Republicii Moldova, Matematica, Number 1 (68), 2012, 90-107.
[6] W. Magnus, F. Oberhettinger and R. P. Soni, formulas and theorems for the special functions of Mathematical Physics, 3rd edition, Springer, 1966
[7] E. Orsingher, Probability law, flow function, maximum distribution of wave-governed random motions and their connections with Kirchoff's laws, Stochastic Processes and their Applications, 34, 49-66, 1990.
[8] A. P. Prudnikov, Y. A. Brychkov and O.I. Marichev, Integrals and Series, Volume 2, Special functions, Gordon and Breach Science Publishers, 1986
[9] A. P. Prudnikov and O. Marichev, Integrals and Series, Volume 4, Direct Laplace transforms, Gordon and Breach Science Publishers, 1992
[10] T. Wakhare, C. Vignat, Q.-N. Le, and S. Robins, A continuous analogue of lattice path enumeration. ArXiV:1707.01616[math.CO].
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